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## **Critical infrastructure operation process related to climate-weather change process including extreme weather hazard**

### **Keywords**

critical infrastructure, operation, prediction, climate-weather change

### **Abstract**

The operation process of the critical infrastructure is considered and its operation states are introduced. The semi-Markov process is used to construct a general probabilistic model of the critical infrastructure operation process. The semi-Markov process is used to construct a general probabilistic model of the climate-weather change process for the critical infrastructure operating area.

### **1. Introduction**

The operation process of the critical infrastructure is considered and its operation states are introduced. The semi-Markov process is used to construct a general probabilistic model of the critical infrastructure operation process. To build this model, the vector of probabilities of the critical infrastructure operation process staying at the initial operation states, the matrix of probabilities of the critical infrastructure operation process transitions between the operation states, the matrix of conditional distribution functions and the matrix of conditional density functions of the critical infrastructure operation process conditional sojourn times at the operation states are defined.

The climate-weather change process for the critical infrastructure operating area is considered and its states are introduced. The semi-Markov process is used to construct a general probabilistic model of the climate-weather change process for the critical infrastructure operating area. To build this model the vector of probabilities of the climate-weather change process staying at the initials climate-weather states, the matrix of probabilities of the climate-weather change process transitions between the climate-weather states, the matrix of conditional distribution

functions and the matrix of conditional density functions of the climate-weather change process conditional sojourn times at the climate-weather states are defined.

Further, these two precesses are joined into a general model of the critical infrastructure operation process related to climate weather change process.

The operation process of a critical infrastructure often has significant influence on its safety. Also, a critical infrastructure operating environment area climate-weather conditions are essential in its safety analysis. Usually, the critical infrastructure operation process and the climate-weather conditions at its operating area interact and have either an explicit or an implicit strong joint influence on the critical infrastructure safety. Thus, considering together those two processes influence on the critical infrastructure safety is of grate practical value.

To construct a joint mdel of those two processes, first, the semi-Markov approaches to a critical infrastructure operation process modeling and to climate-weather change process are separately developed. Next, those two separate models are linked into a jeneral joint model of a critical infrastructure operation process related to the climate-weather change process including extreme hazards. is build. climate to modelling a critical

infrastructure operation process including operating environment threats.

The method of defining the parameters of this general joint model of the critical infrastructure operation process is presented. The procedures of these parameters determination using the parameters of the two separate models and their application to the critical infrastructure operation characteristics related to climate-weather change prediction are proposed as well.

## 2. Critical infrastructure operation process-modelling

### 2.1. Semi-Markov model of critical infrastructure operation process

We assume that the critical infrastructure during its operation process is taking  $\nu, \nu \in N$ , different operation states  $z_1, z_2, \dots, z_\nu$ . Further, we define the critical infrastructure operation process  $Z(t)$ ,  $t \in (-\infty, +\infty)$ , with discrete operation states from the set  $\{z_1, z_2, \dots, z_\nu\}$ . Moreover, we assume that the critical infrastructure operation process  $Z(t)$  is a semi-Markov process [Grabski, 2002], [Limnios, 2005], [Mercier, 2008], [Soszyńska, 2007], [Kołowrocki, Soszyńska, 2011] with the conditional sojourn times  $\theta_{bl}$  at the operation states  $z_b$  when its next operation state is  $z_l$ ,  $b, l = 1, 2, \dots, \nu$ ,  $b \neq l$ . Under these assumptions, the critical infrastructure operation process may be described by:

- the vector of the initial probabilities  $p_b(0) = P(Z(0) = z_b)$ ,  $b = 1, 2, \dots, \nu$ , of the critical infrastructure operation process  $Z(t)$  staying at particular operation states at the moment  $t = 0$

$$[p_b(0)]_{1 \times \nu} = [p_1(0), p_2(0), \dots, p_\nu(0)]; \quad (1)$$

- the matrix of probabilities  $p_{bl}$ ,  $b, l = 1, 2, \dots, \nu$ , of the critical infrastructure operation process  $Z(t)$  transitions between the operation states  $z_b$  and  $z_l$

$$[p_{bl}]_{\nu \times \nu} = \begin{bmatrix} p_{11} & p_{12} & \dots & p_{1\nu} \\ p_{21} & p_{22} & \dots & p_{2\nu} \\ \dots & & & \\ p_{\nu 1} & p_{\nu 2} & \dots & p_{\nu \nu} \end{bmatrix}, \quad (2)$$

where by formal agreement

$$p_{bb} = 0 \text{ for } b = 1, 2, \dots, \nu;$$

- the matrix of conditional distribution functions  $H_{bl}(t) = P(\theta_{bl} < t)$ ,  $b, l = 1, 2, \dots, \nu$ , of the critical infrastructure operation process  $Z(t)$  conditional sojourn times  $\theta_{bl}$  at the operation states

$$[H_{bl}(t)]_{\nu \times \nu} = \begin{bmatrix} H_{11}(t) & H_{12}(t) & \dots & H_{1\nu}(t) \\ H_{21}(t) & H_{22}(t) & \dots & H_{2\nu}(t) \\ \dots & & & \\ H_{\nu 1}(t) & H_{\nu 2}(t) & \dots & H_{\nu \nu}(t) \end{bmatrix}, \quad (3)$$

where by formal agreement

$$H_{bb}(t) = 0 \text{ for } b = 1, 2, \dots, \nu,$$

We assume that the suitable and typical distributions suitable to describe the critical infrastructure operation process  $Z(t)$  conditional sojourn times  $\theta_{bl}$ ,  $b, l = 1, 2, \dots, \nu$ ,  $b \neq l$ , in the particular operation states are that defined in [Kołowrocki, Soszyńska 2011] and [EU-CIRCLE Report D2.1-GMU2, 2016].

## 3. Critical infrastructure operation process – prediction

### 3.1. Prediction of critical infrastructure operation process characteristics

Assuming that we have identified the unknown parameters of the critical infrastructure operation process semi-Markov model:

– the initial probabilities  $p_b(0)$ ,  $b = 1, 2, \dots, \nu$ , of the critical infrastructure operation process staying at the particular state  $z_b$  at the moment  $t = 0$ ;

– the probabilities  $p_{bl}$ ,  $b, l = 1, 2, \dots, \nu$ ,  $b \neq l$ , of the critical infrastructure operation process transitions from the operation state  $z_b$  into the operation state  $z_l$ ;

– the distributions of the critical infrastructure operation process conditional sojourn times  $\theta_{bl}$ ,  $b, l = 1, 2, \dots, \nu$ ,  $b \neq l$ , at the particular operation states and their mean values  $M_{bl} = E[\theta_{bl}]$ ,  $b, l = 1, 2, \dots, \nu$ ,  $b \neq l$ ;

we can predict this process basic characteristics.

As the mean values of the conditional sojourn times  $\theta_{bl}$  are given by [Kołowrocki, Soszyńska-Budny, 2011]

$$M_{bl} = E[\theta_{bl}] = \int_0^\infty t dH_{bl}(t) = \int_0^\infty t h_{bl}(t) dt, \quad (6)$$

$$b, l = 1, 2, \dots, \nu, \quad b \neq l,$$

then for the distinguished distributions (2.5)-(2.11) in [EU-CRCLE Report D2.1-GMU2, 2016], the mean values of the system operation process  $Z(t)$  conditional sojourn times  $\theta_{bl}$ ,  $b, l = 1, 2, \dots, y$ ,  $b \neq l$ , at the particular operation states are respectively given by (2.6)-(2.12) in [EU-CRCLE Report D2.1-GMU2, 2016].

From the formula for total probability, it follows that the unconditional distribution functions of the sojourn times  $\theta_b$ ,  $b = 1, 2, \dots, y$ , of the system operation process  $Z(t)$  at the operation states  $z_b$ ,  $b = 1, 2, \dots, y$ , are given by [Kołowrocki, Soszyńska-Budny, 2011]

$$H_b(t) = \sum_{l=1}^y p_{bl} H_{bl}(t), \quad b = 1, 2, \dots, y. \quad (7)$$

Hence, the mean values  $E[\theta_b]$  of the system operation process  $Z(t)$  unconditional sojourn times  $\theta_b$ ,  $b = 1, 2, \dots, y$ , at the operation states are given by

$$M_b = E(\theta_b) = \sum_{l=1}^y p_{bl} M_{bl}, \quad b = 1, 2, \dots, y, \quad (8)$$

where  $M_{bl}$  are defined by the formula (6) in a case of any distribution of sojourn times  $\theta_{bl}$  and by the formulae (2.6)-(2.12) in the cases of particular defined respectively by (2.5)-(2.11) in [EU-CRCLE Report D2.1-GMU2, 2016] distributions of these sojourn times.

The limit values of the system operation process  $Z(t)$  transient probabilities at the particular operation states

$$p_b(t) = P(Z(t) = z_b), \quad t \in \langle 0, +\infty \rangle, \quad b = 1, 2, \dots, y, \quad (9)$$

are given by [Kołowrocki, Soszyńska-Budny, 2011]

$$p_b = \lim_{t \rightarrow \infty} p_b(t) = \frac{\pi_b M_b}{\sum_{l=1}^y \pi_l M_l}, \quad b = 1, 2, \dots, y, \quad (10)$$

where  $M_b$ ,  $b = 1, 2, \dots, y$ , are given by (8), while the steady probabilities  $\pi_b$  of the vector  $[\pi_b]_{1 \times y}$  satisfy the system of equations

$$\begin{cases} [\pi_b] = [\pi_b][p_{bl}] \\ \sum_{l=1}^y \pi_l = 1. \end{cases} \quad (11)$$

In the case of a periodic system operation process, the limit transient probabilities  $p_b$ ,  $b = 1, 2, \dots, y$ , at the

operation states defined by (10), are the long term proportions of the system operation process  $Z(t)$  sojourn times at the particular operation states  $z_b$ ,  $b = 1, 2, \dots, y$ .

Other interesting characteristics of the system operation process  $Z(t)$  possible to obtain are its total sojourn times  $\hat{\theta}_b$  at the particular operation states  $z_b$ ,  $b = 1, 2, \dots, y$ , during the fixed system operation time. It is well known [Kołowrocki, Soszyńska-Budny, 2011] that the system operation process total sojourn times  $\hat{\theta}_b$  at the particular operation states  $z_b$ , for sufficiently large operation time  $\theta$ , have approximately normal distributions with the expected value given by

$$\hat{M}_b = E[\hat{\theta}_b] = p_b \theta, \quad b = 1, 2, \dots, y, \quad (12)$$

where  $p_b$  are given by (10).

## 4. Climate-weather change process - modelling

### 4.1. Semi-Markov model of climate-weather change process

To model the climate-weather change process for the critical infrastructure operating area we assume that the climate-weather in this area is taking  $w$ ,  $w \in N$ , different climate-weather states  $c_1, c_2, \dots, c_w$ . Further, we define the climate-weather change process  $C(t)$ ,  $t \in \langle 0, +\infty \rangle$ , with discrete operation states from the set  $\{c_1, c_2, \dots, c_w\}$ . Assuming that the climate-weather change process  $C(t)$  is a semi-Markov process it can be described by:

- the number of climate-weather states  $w$ ,  $w \in N$ ,
- the vector

$$[q_b(0)]_{1 \times w} = [q_1(0), q_2(0), \dots, q_w(0)]$$

of the initial probabilities

$$q_b(0) = P(C(0) = c_b), \quad b = 1, 2, \dots, w,$$

of the climate-weather change process  $C(t)$  staying at particular climate-weather states  $c_b$  at the moment  $t = 0$ ;

- the matrix

$$[q_{bl}]_{w \times w} = \begin{bmatrix} q_{11} & q_{12} & \dots & q_{1w} \\ q_{21} & q_{22} & \dots & q_{2w} \\ \dots & \dots & \dots & \dots \\ q_{w1} & q_{w2} & \dots & q_{ww} \end{bmatrix} \quad (14)$$

of the probabilities of transitions  $q_{bl}$ ,  $b, l = 1, 2, \dots, w$ ,  $b \neq l$ , of the climate-weather change process  $C(t)$  from the climate-weather states  $c_b$  to  $c_l$ , where by formal agreement

$$q_{bb} = 0 \text{ for } b = 1, 2, \dots, w;$$

– the matrix

$$[C_{bl}(t)]_{wxw} = \begin{bmatrix} C_{11}(t) C_{12}(t) \dots C_{1w}(t) \\ C_{21}(t) C_{22}(t) \dots C_{2w}(t) \\ \dots \\ C_{w1}(t) C_{w2}(t) \dots C_{ww}(t) \end{bmatrix}$$

of the conditional distribution functions

$$C_{bl}(t) = P(C_{bl} < t), \quad b, l = 1, 2, \dots, w,$$

of the conditional sojourn times  $C_{bl}$  at the climate-weather states  $c_b$  when its next climate-weather state is

$c_l$ ,  $b, l = 1, 2, \dots, w$ ,  $b \neq l$ , where by formal agreement

$$C_{bb}(t) = 0 \text{ for } b = 1, 2, \dots, w,$$

or equivalently the matrix

$$[c_{bl}(t)]_{wxw} = \begin{bmatrix} c_{11}(t) c_{12}(t) \dots c_{1w}(t) \\ c_{21}(t) c_{22}(t) \dots c_{2w}(t) \\ \dots \\ c_{w1}(t) c_{w2}(t) \dots c_{ww}(t) \end{bmatrix} \quad (16)$$

of the conditional density functions of the climate-weather change process  $C(t)$  conditional sojourn times  $C_{bl}$  at the climate-weather states corresponding to the conditional distribution functions  $C_{bl}(t)$ , where

$$c_{bl}(t) = \frac{d}{dt}[C_{bl}(t)] \text{ for } b, l = 1, 2, \dots, w, \quad b \neq l, \quad (17)$$

and by formal agreement

$$c_{bb}(t) = 0 \text{ for } b = 1, 2, \dots, w.$$

We assume that the suitable and typical distributions suitable to describe the climate-weather change process  $C(t)$  conditional sojourn times  $C_{bl}$ ,  $b, l = 1, 2, \dots, w$ ,  $b \neq l$ , at the particular climate-weather states are given by (4.5)-(4.12) in [EU-CRCLE Report D2.1-GMU2, 2016]

## 5. Climate-weather change process – prediction

### 5.1. Prediction of climate-weather process characteristics

Assuming that we have identified the unknown parameters of the climate-weather change process semi-Markov model:

– the initial probabilities  $q_b(0)$ ,  $b = 1, 2, \dots, w$ , of the climate-weather change process staying at the particular state  $c_b$  at the moment  $t = 0$ ;

– the probabilities  $q_{bl}$ ,  $b, l = 1, 2, \dots, w$ ,  $b \neq l$ , of the climate-weather change process transitions from the climate-weather state  $c_b$  into the climate-weather state  $c_l$ ;

– the distributions of the climate-weather change process conditional sojourn times  $C_{bl}$ ,  $b, l = 1, 2, \dots, w$ ,  $b \neq l$ , at the particular climate-weather states and their mean values  $M_{bl} = E[C_{bl}]$ ,  $b, l = 1, 2, \dots, w$ ; ; we can predict this process basic characteristics.

As the mean values of the conditional sojourn times  $C_{bl}$  are given by [Kołowrocki, Soszyńska-Budny, 2011]

$$N_{bl} = E[C_{bl}] = \int_0^{\infty} t dC_{bl}(t) = \int_0^{\infty} t c_{bl}(t) dt, \quad (18)$$

$$b, l = 1, 2, \dots, w, \quad b \neq l,$$

then for the distinguished distributions (4.5)-(4.12) in [Kołowrocki, Soszyńska-Budny 2011], the mean values of the climate-weather change process  $C(t)$  conditional sojourn times  $C_{bl}$ ,  $b, l = 1, 2, \dots, w$ ,  $b \neq l$ , at the particular operation states are respectively given by (4.14)-(4.21) [Kołowrocki, Soszyńska-Budny 2011].

From the formula for total probability, it follows that the unconditional distribution functions of the sojourn times  $C_b$ ,  $b = 1, 2, \dots, w$ , of the climate-weather change process  $C(t)$  at the climate-weather states  $c_b$ ,  $b = 1, 2, \dots, w$ , are given by [Kołowrocki, Soszyńska-Budny, 2011]

$$C_b(t) = \sum_{l=1}^w q_{bl} C_{bl}(t), \quad b = 1, 2, \dots, w.$$

Hence, the mean values  $E[C_b]$  of the climate-weather change process  $C(t)$  unconditional sojourn times  $C_b$ ,  $b = 1, 2, \dots, w$ , at the climate-weather states are given by

$$N_b = E[C_b] = \sum_{l=1}^v q_{bl} N_{bl}, \quad b = 1, 2, \dots, w, \quad (20)$$

where  $N_{bl}$  are defined by the formula (18) in a case of any distribution of sojourn times  $C_{bl}$  and by the formulae (4.14)-(4.21) in [EU-CRCLE Report D2.1-GMU2, 2016] in the cases of particular defined respectively by (4.5)-(4.12) [EU-CRCLE Report D2.1-GMU2, 2016] distributions of these sojourn times.

The limit values of the climate-weather change process  $C(t)$  transient probabilities at the particular operation states

$$q_b(t) = P(C(t) = c_b), \quad t \in \langle 0, +\infty \rangle, \quad b = 1, 2, \dots, w, \quad (21)$$

are given by [Kołowrocki, Soszyńska-Budny, 2016]

$$q_b = \lim_{t \rightarrow \infty} q_b(t) = \frac{\pi_b N_b}{\sum_{l=1}^w \pi_l N_l}, \quad b = 1, 2, \dots, w, \quad (22)$$

where  $N_b$ ,  $b = 1, 2, \dots, w$ , are given by (20), while the steady probabilities  $\pi_b$  of the vector  $[\pi_b]_{1 \times w}$  satisfy the system of equations

$$\begin{cases} [\pi_b] = [\pi_b][q_{bl}] \\ \sum_{l=1}^w \pi_l = 1. \end{cases} \quad (23)$$

In the case of a periodic climate-weather change process, the limit transient probabilities  $q_b$ ,  $b = 1, 2, \dots, w$ , at the climate-weather states defined by (22), are the long term proportions of the climate-weather change process  $C(t)$  sojourn times at the particular climate-weather states  $c_b$ ,  $b = 1, 2, \dots, w$ .

Other interesting characteristics of the system climate-weather change process  $C(t)$  possible to obtain are its total sojourn times  $\hat{C}_b$  at the particular climate-weather states  $c_b$ ,  $b = 1, 2, \dots, w$ , during the fixed time. It is well known [Kołowrocki, Soszyńska-Budny, 2011] that the climate-weather change process total sojourn times  $\hat{C}_b$  at the particular climate-weather states  $c_b$ , for sufficiently large time  $\theta$ , have approximately normal distributions with the expected value given by

$$\hat{N}_b = E[\hat{C}_b] = q_b \theta, \quad b = 1, 2, \dots, w, \quad (24)$$

where  $q_b$  are given by (22).

## 6. Critical infrastructure operation process related to climate-weather change process-modelling

We assume, that the critical infrastructure during its operation process is taking  $v$ ,  $v \in N$ , different operation states  $z_1, z_2, \dots, z_v$ . Further, we define the critical infrastructure operation process  $Z(t)$ ,  $t \in \langle 0, +\infty \rangle$ , with discrete operation states from the set  $\{z_1, z_2, \dots, z_v\}$ . Moreover, we assume that the critical infrastructure operation process  $Z(t)$  is a semi-Markov process that can be described by:

- the vector  $[p_b(0)]_{1 \times v}$  of the initial probabilities  $p_b(0)$ ,  $b = 1, 2, \dots, v$ , of the critical infrastructure operation process  $Z(t)$  staying at particular operation state  $z_b$ ,  $b = 1, 2, \dots, v$ , at the moment  $t = 0$ ;

- the matrix  $[p_{bl}]_{v \times v}$  of probabilities  $p_{bl}$ ,  $b, l = 1, 2, \dots, v$ , of the critical infrastructure operation process  $Z(t)$  transitions between the operation states  $z_b$  and  $z_l$ ,  $b, l = 1, 2, \dots, v$ ;

- the matrix  $[H_{bl}(t)]_{v \times v}$  of conditional distribution functions  $[H_{bl}(t)]$ ,  $b, l = 1, 2, \dots, v$ , of the critical infrastructure operation process  $Z(t)$  conditional sojourn times  $\theta_{bl}$  at the operation states  $z_b$  under the condition that the next operation state will be  $z_l$ ,  $b, l = 1, 2, \dots, v$ .

We assume that the climate-weather change process  $C(t)$ ,  $t \in \langle 0, +\infty \rangle$ , at the critical infrastructure operating area is taking  $w$ ,  $w \in N$ , different climate-weather states  $c_1, c_2, \dots, c_w$ . Further, we assume that the climate-weather change process  $C(t)$  is a semi-Markov process and it can be described by:

- the vector  $[q_b(0)]_{1 \times w}$  of the initial probabilities  $q_b(0)$ ,  $b = 1, 2, \dots, w$ , of the climate-weather change process  $C(t)$  staying at particular climate-weather states  $c_b$ ,  $b = 1, 2, \dots, w$ , at the moment  $t = 0$ ;

- the matrix  $[q_{bl}]_{w \times w}$  of the probabilities  $q_{bl}$ ,  $b, l = 1, 2, \dots, w$ , of transitions of the climate-weather change process  $C(t)$  from the climate-weather states  $c_b$  to the climate-weather state  $c_l$ ,  $b, l = 1, 2, \dots, w$ ;

- the matrix  $[C_{bl}(t)]_{w \times w}$  of the conditional distribution functions  $C_{bl}(t)$ ,  $b, l = 1, 2, \dots, w$ , of the conditional sojourn times  $C_{bl}$  at the climate-weather states  $c_b$  when its next climate-weather state is  $c_l$ ,  $b, l = 1, 2, \dots, w$ .

### 6.1. Joint model of independent critical infrastructure operation process and climate-weather change process

Under the assumption that the critical infrastructure operation process  $Z(t)$ ,  $t \in \langle 0, +\infty \rangle$ , and the climate-weather change process  $C(t)$  are independent, we introduce the joint process of critical infrastructure operation process and climate-weather change process called the critical infrastructure operation process related to climate-weather change marked by

$$ZC(t), t \in \langle 0, +\infty \rangle, \quad (25)$$

and we assume that it can take  $v, w, v, w \in N$ , different operation states

$$z_{c_{11}}, z_{c_{12}}, \dots, z_{c_{vw}}, \quad i = 1, 2, \dots, v, \quad j = 1, 2, \dots, w, \quad (26)$$

We assume that the critical infrastructure operation process related to climate-weather change  $ZC(t)$ , at the moment  $t \in \langle 0, +\infty \rangle$ , is at the state  $z_{c_{ij}}$ ,  $i = 1, 2, \dots, v$ ,  $j = 1, 2, \dots, w$ , if and only if at that moment, the operation process  $Z(t)$  is at the operation states  $z_i$ ,  $i = 1, 2, \dots, v$ , and the climate-weather change process  $C(t)$  is at the climate-weather state  $c_j$ ,  $j = 1, 2, \dots, w$ , what we mark as follows:

$$(ZC(t) = z_{c_{ij}}) \Leftrightarrow (Z(t) = z_i \cap C(t) = c_j), \quad (27)$$

$$t \in \langle 0, +\infty \rangle, i = 1, 2, \dots, v, j = 1, 2, \dots, w.$$

Further, we define the initial probabilities

$$pq_{ij}(0) = P(ZC(0) = z_{c_{ij}}), i = 1, 2, \dots, v, \quad (28)$$

$$j = 1, 2, \dots, w,$$

of the critical infrastructure operation process related to climate-weather change  $ZC(t)$ , at the initial moment  $t = 0$  at the operation and climate-weather state  $z_{c_{ij}}$ ,  $i = 1, 2, \dots, v \in N$ ,  $j = 1, 2, \dots, w$ , and this way we have the vector

$$[pq_{ij}(0)]_{1 \times vw}$$

$$= \begin{bmatrix} pq_{11}(0), pq_{12}(0), \dots, pq_{1w}(0); pq_{21}(0), \\ pq_{22}(0), \dots, pq_{2w}(0); \dots; pq_{v1}(0), \\ pq_{v2}(0), \dots, pq_{vw}(0) \end{bmatrix} \quad (29)$$

of the initial probabilities the critical infrastructure operation process related to climate-weather change  $ZC(t)$  staying at the particular operation and climate-weather state at the initial moment  $t = 0$ .

From the assumption that the critical infrastructure operation process  $Z(t)$  and climate-weather change process  $C(t)$  are independent, it follows that

$$pq_{ij}(0) = P(ZC(0) = z_{c_{ij}})$$

$$= P(Z(0) = z_i \cap C(0) = c_j)$$

$$= P(Z(0) = z_i) \cdot P(C(0) = c_j) = p_i(0) \cdot q_j(0), \quad (30)$$

$$i = 1, 2, \dots, v, j = 1, 2, \dots, w,$$

where  $p_i(0)$ ,  $i = 1, 2, \dots, v$ , and  $q_j(0)$ ,  $j = 1, 2, \dots, w$ , are respectively defined in Section 2.1 and Section 4.1. [EU-CIRCLE Report D2.1-GMU2, 2016].

Hence, the vector of the initial probabilities the critical infrastructure operation process related to climate-weather change  $ZC(t)$  defined by (5.29) takes the following form

$$[pq_{ij}(0)]_{1 \times vw} = [p_i(0)q_j(0)]_{1 \times vw}$$

$$= \begin{bmatrix} p_1(0)q_1(0), p_1(0)q_2(0), \dots, p_1(0) \\ q_{1w}(0); \dots; p_v(0)q_1(0), p_v(0) \\ q_2(0), \dots, p_v(0)q_w(0) \end{bmatrix}. \quad (31)$$

Further, we introduce the probabilities

$$pq_{ijkl}, \quad i = 1, 2, \dots, v, j = 1, 2, \dots, w, \quad (32)$$

$$k = 1, 2, \dots, v, l = 1, 2, \dots, w,$$

of the transitions of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  between the operation states

$$z_{c_{ij}} \text{ and } z_{c_{kl}}, \quad i = 1, 2, \dots, v, \quad (33)$$

$$j = 1, 2, \dots, w, k = 1, 2, \dots, v, l = 1, 2, \dots, w,$$

and get their following matrix form



$$[pq_{ijkl}]_{w \times w} = \begin{bmatrix} pq_{1111} & pq_{1112} & \dots & pq_{111w} & pq_{1121} & pq_{1122} & \dots & pq_{112w} & \dots & pq_{11v1} & pq_{11v2} & \dots & pq_{11vw} \\ pq_{1211} & pq_{1212} & \dots & pq_{121w} & pq_{1221} & pq_{1222} & \dots & pq_{122w} & \dots & pq_{12v1} & pq_{12v2} & \dots & pq_{12vw} \\ \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots \\ pq_{wv11} & pq_{wv12} & \dots & pq_{wv1w} & pq_{wv21} & pq_{wv22} & \dots & pq_{wv2w} & \dots & pq_{wvv1} & pq_{wvv2} & \dots & pq_{wvww} \end{bmatrix}. \quad (34)$$

From the assumption that the critical infrastructure operation process  $Z(t)$  and climate-weather change process  $C(t)$  are independent, it follows that

$$pq_{ijkl} = p_{ik}q_{jl}, \quad i = 1, 2, \dots, v, \quad j = 1, 2, \dots, w, \quad k = 1, 2, \dots, v, \quad l = 1, 2, \dots, w, \quad (35)$$

where

$$p_{ik}, \quad i = 1, 2, \dots, v, \quad k = 1, 2, \dots, v, \quad \text{and} \quad q_{jl}, \quad j = 1, 2, \dots, w, \quad l = 1, 2, \dots, w, \quad (36)$$

are respectively defined in Section 2.1 and Section 4.1 [EU-CIRCLE Report D2.1-GMU2, 2016].

Hence, the matrix of the probabilities of transitions between the critical infrastructure operation process related to climate-weather change  $ZC(t)$  defined by (24) takes the following form

$$[pq_{ijkl}]_{w \times w} = [p_{ik}q_{jl}]_{w \times w} = \begin{bmatrix} p_{11}q_{11} & p_{11}q_{12} & \dots & p_{11}q_{1w} & p_{12}q_{11} & p_{12}q_{12} & \dots & p_{12}q_{1w} & \dots & p_{1v}q_{11} & p_{1v}q_{12} & \dots & p_{1v}q_{1w} \\ p_{11}q_{21} & p_{11}q_{22} & \dots & p_{11}q_{2w} & p_{12}q_{21} & p_{12}q_{22} & \dots & p_{12}q_{2w} & \dots & p_{1v}q_{21} & p_{1v}q_{22} & \dots & p_{1v}q_{2w} \\ \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots \\ p_{v1}q_{w1} & p_{v1}q_{w2} & \dots & p_{v1}q_{ww} & p_{v2}q_{w1} & p_{v2}q_{w2} & \dots & p_{v2}q_{ww} & \dots & p_{vv}q_{w1} & p_{vv}q_{w2} & \dots & p_{vv}q_{ww} \end{bmatrix}. \quad (37)$$

The matrix of conditional distribution functions

$$HC_{ijkl}(t) = P(\theta C_{ijkl} < t), \quad t \in \langle 0, +\infty \rangle, \quad i = 1, 2, \dots, v, \quad j = 1, 2, \dots, w, \quad k = 1, 2, \dots, v, \quad l = 1, 2, \dots, w, \quad (38)$$

of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  conditional sojourn times  $\theta C_{ijkl}$ ,  $i = 1, 2, \dots, v$ ,  $j = 1, 2, \dots, w$ ,  $k = 1, 2, \dots, v$ ,  $l = 1, 2, \dots, w$ , at the operation state  $zc_{ik}$ ,  $i = 1, 2, \dots, v$ ,  $k = 1, 2, \dots, v$ , when the next operation state is  $zc_{jl}$ ,  $j = 1, 2, \dots, w$ ,  $l = 1, 2, \dots, w$ , takes the following form

$$[HC_{ijkl}(t)]_{w \times w} = \begin{bmatrix} HC_{1111}(t) & HC_{1112}(t) & \dots & HC_{111w}(t); & HC_{1121}(t) & HC_{1122}(t) & \dots & HC_{112w}(t); & \dots; & HC_{11v1}(t) & HC_{11v2}(t) & \dots & HC_{11vw}(t) \\ HC_{1211}(t) & HC_{1212}(t) & \dots & HC_{121w}(t); & HC_{1221}(t) & HC_{1222}(t) & \dots & HC_{122w}(t); & \dots; & HC_{12v1}(t) & HC_{12v2}(t) & \dots & HC_{12vw}(t) \\ \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots \\ HC_{wv11}(t) & HC_{wv12}(t) & \dots & HC_{wv1w}(t); & HC_{wv21}(t) & HC_{wv22}(t) & \dots & HC_{wv2w}(t); & \dots; & HC_{wvv1}(t) & HC_{wvv2}(t) & \dots & HC_{wvww}(t) \end{bmatrix} \quad (39)$$

and the matrix of their corresponding conditional density functions

$$hc_{ijkl}(t) = \frac{d}{dt} [HC_{ijkl}(t)] \quad \text{for } t \in \langle 0, +\infty \rangle, \quad i = 1, 2, \dots, v, \quad j = 1, 2, \dots, w, \quad k = 1, 2, \dots, v, \quad l = 1, 2, \dots, w, \quad (40)$$



the form

$$\begin{aligned}
 & [hc_{ij\,kl}(t)]_{\text{maxvw}} \\
 & = \begin{bmatrix} hc_{1111}(t) hc_{1112}(t) \dots hc_{111w}(t); hc_{1121}(t) hc_{1122}(t) \dots hc_{112w}(t); \dots; hc_{11v1}(t) hc_{11v2}(t) \dots hc_{11vw}(t) \\ hc_{1211}(t) hc_{1212}(t) \dots hc_{121w}(t); hc_{1221}(t) hc_{1222}(t) \dots hc_{122w}(t); \dots; hc_{12v1}(t) hc_{12v2}(t) \dots hc_{12vw}(t) \\ \dots \\ hc_{vw11}(t) hc_{vw12}(t) \dots hc_{vw1w}(t); hc_{vw21}(t) hc_{vw22}(t) \dots hc_{vw2w}(t); \dots; hc_{wv11}(t) hc_{wv12}(t) \dots hc_{wv1w}(t) \end{bmatrix}. \quad (41)
 \end{aligned}$$

From the assumption that the critical infrastructure operation process  $Z(t)$  and climate-weather change process  $C(t)$  are independent, it follows that

$$\begin{aligned}
 HC_{ij\,kl}(t) &= P(\theta C_{ij\,kl} < t) = P(\theta_{ik} < t \cap C_{jl} < t) = H_{ik}(t)C_{jl}(t), \quad t \in \langle 0, +\infty \rangle, \\
 i &= 1, 2, \dots, v, \quad j = 1, 2, \dots, w, \quad k = 1, 2, \dots, v, \quad l = 1, 2, \dots, w, \quad (42)
 \end{aligned}$$

and

$$\begin{aligned}
 hc_{ij\,kl}(t) &= \frac{d}{dt}[HC_{ij\,kl}(t)] = \frac{d}{dt}[H_{ik}(t)C_{jl}(t)] = h_{ik}(t)C_{jl}(t) + H_{ik}(t)c_{jl}(t), \quad t \in \langle 0, +\infty \rangle, \\
 i &= 1, 2, \dots, v, \quad j = 1, 2, \dots, w, \quad k = 1, 2, \dots, v, \quad l = 1, 2, \dots, w, \quad (43)
 \end{aligned}$$

where

$$H_{ik}(t), \quad i = 1, 2, \dots, v, \quad k = 1, 2, \dots, v, \quad \text{and} \quad C_{jl}(t), \quad j = 1, 2, \dots, w, \quad l = 1, 2, \dots, w, \quad (44)$$

and

$$h_{ik}(t), \quad i = 1, 2, \dots, v, \quad k = 1, 2, \dots, v, \quad \text{and} \quad c_{jl}(t), \quad j = 1, 2, \dots, w, \quad l = 1, 2, \dots, w, \quad (45)$$

are respectively defined in Chapter 2, Section 2.1 and Chapter 4, Section 4.1 [EU-CIRCLE Report D2.1-GMU2, 2016].

Hence, the matrix of the conditional distribution functions and the matrix of the conditional density functions of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  conditional sojourn times defined by (29) and (31) respectively take the following forms

$$\begin{aligned}
 & [HC_{ij\,kl}(t)]_{\text{maxvw}} = [H_{ik}(t)C_{jl}(t)]_{\text{maxvw}} \\
 & = \begin{bmatrix} H_{11}(t)C_{11}(t) H_{11}(t)C_{12}(t) \dots H_{11}(t)C_{1w}(t); \dots; H_{1v}(t)C_{11}(t) H_{1v}(t)C_{12}(t) \dots H_{1v}(t)C_{1w}(t) \\ H_{11}(t)C_{21}(t) H_{11}(t)C_{22}(t) \dots H_{11}(t)C_{2w}(t); \dots; H_{1v}(t)C_{21}(t) H_{1v}(t)C_{22}(t) \dots H_{1v}(t)C_{2w}(t) \\ \dots \\ H_{v1}(t)C_{w1}(t) H_{v1}(t)C_{w2}(t) \dots H_{v1}(t)C_{vw}(t); \dots; H_{vv}(t)C_{w1}(t) H_{vv}(t)C_{w2}(t) \dots H_{vv}(t)C_{vw}(t) \end{bmatrix} \quad (46)
 \end{aligned}$$

and

$$[hc_{ij\,kl}(t)]_{\text{maxvw}} = [h_{ik}(t)C_{jl}(t) + H_{ik}(t)c_{jl}(t)]_{\text{maxvw}}$$



$$= \left[ \begin{array}{l} h_{11}(t)C_{11}(t) + H_{11}(t)c_{11}(t) \dots h_{11}(t)C_{1w}(t) + H_{11}(t)c_{1w}(t); \dots; h_{1v}(t)C_{11}(t) + H_{1v}(t)c_{11}(t) \dots h_{1v}(t)C_{1w}(t) + H_{1v}(t)c_{1w}(t) \\ h_{11}(t)C_{21}(t) + H_{11}(t)c_{21}(t) \dots h_{11}(t)C_{2w}(t) + H_{11}(t)c_{2w}(t); \dots; h_{1v}(t)C_{21}(t) + H_{1v}(t)c_{21}(t) \dots h_{1v}(t)C_{2w}(t) + H_{1v}(t)c_{2w}(t) \\ \dots \\ h_{v1}(t)C_{w1}(t) + H_{v1}(t)c_{w1}(t) \dots h_{v1}(t)C_{vw}(t) + H_{v1}(t)c_{vw}(t); \dots; h_{vv}(t)C_{w1}(t) + H_{vv}(t)c_{w1}(t) \dots h_{vv}(t)C_{vw}(t) + H_{vv}(t)c_{vw}(t) \end{array} \right] \quad (47)$$

We assume that the suitable and typical distributions suitable to describe the critical infrastructure operation process  $Z(t)$  conditional sojourn times  $\theta_{bl}$ ,  $b, l = 1, 2, \dots, \nu$ ,  $b \neq l$ , in the particular operation states are that defined in [Kołowrocki, Soszyńska-Budny, 2011], [EU-CIRCLE Report D2.1-GMU4-Part1, 2016] and [EU-CIRCLE Report D2.1-GMU4-Part2, 2016].

### 6.2. Joint model of dependent critical infrastructure operation process and climate-weather change process

Under the assumption that the critical infrastructure operation process  $Z(t)$ ,  $t \in \langle 0, +\infty \rangle$ , and the climate-weather change process  $C(t)$  are dependent, we introduce the joint process of critical infrastructure operation process and climate-weather change process called the critical infrastructure operation process related to climate-weather change marked by

$$ZC(t), \quad t \in \langle 0, +\infty \rangle, \quad (48)$$

and we assume that it can take  $\nu w$ ,  $\nu, w \in N$ , different operation states

$$z_{c_{11}}, z_{c_{12}}, \dots, z_{c_{\nu w}}, \quad i = 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w, \quad (49)$$

We assume that the critical infrastructure operation process related to climate-weather change  $ZC(t)$ , at the moment  $t \in \langle 0, +\infty \rangle$ , is at the state  $z_{c_{ij}}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ , if and only if at that moment, the operation process  $Z(t)$  is at the operation states  $z_i$ ,  $i = 1, 2, \dots, \nu$ , and the climate-weather change process  $C(t)$  is at the climate-weather state  $c_j$ ,  $j = 1, 2, \dots, w$ , what we mark as follows:

$$(ZC(t) = z_{c_{ij}}) \Leftrightarrow (Z(t) = z_i \cap C(t) = c_j), \quad (50)$$

$$t \in \langle 0, +\infty \rangle, \quad i = 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w.$$

Further, we define the initial probabilities

$$pq_{ij}(0) = P(ZC(0) = z_{c_{ij}}), \quad i = 1, 2, \dots, \nu, \quad (51)$$

$$j = 1, 2, \dots, w,$$

of the critical infrastructure operation process related to climate-weather change  $ZC(t)$ , at the initial

moment  $t = 0$  at the operation and climate-weather state  $z_{c_{ij}}$ ,  $i = 1, 2, \dots, \nu \in N$ ,  $j = 1, 2, \dots, w$ , and this way we have the vector

$$[pq_{ij}(0)]_{i \times \nu w}$$

$$= \left[ \begin{array}{l} pq_{11}(0), pq_{12}(0), \dots, pq_{1w}(0); \\ pq_{21}(0), pq_{22}(0), \dots, pq_{2w}(0); \\ \dots; pq_{\nu 1}(0), pq_{\nu 2}(0), \dots, pq_{\nu w}(0) \end{array} \right] \quad (52)$$

of the initial probabilities the critical infrastructure operation process related to climate-weather change  $ZC(t)$  staying at the particular operation and climate-weather state at the initial moment  $t = 0$ . In the case when the processes  $Z(t)$  and  $C(t)$  are dependent the initial probabilities existing in (52) can be expressed either by

$$pq_{ij}(0) = P(ZC(0) = z_{c_{ij}})$$

$$= P(Z(0) = z_i \cap C(0) = c_j)$$

$$= P(Z(0) = z_i) \cdot P(C(0) = c_j \mid Z(0) = z_i)$$

$$= p_i(0) \cdot q_{j|i}(0), \quad i = 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w, \quad (53)$$

where

$$p_i(0) = P(Z(0) = z_i), \quad i = 1, 2, \dots, \nu, \quad (54)$$

are the initial probabilities of the operation process  $Z(t)$  defined in Chapter 2 and

$$q_{j|i}(0) = P(C(0) = c_j \mid Z(0) = z_i), \quad (55)$$

$$i = 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w,$$

are conditional initial probabilities of the climate-weather change process  $C(t)$  defined in Chapter 4 in case they are not conditional or by

$$pq_{ij}(0) = P(ZC(0) = z_{c_{ij}}) = P(Z(0) = z_i \cap C(0) = c_j)$$

$$\begin{aligned}
 &= P(C(0) = C_j) \cdot P(Z(0) = z_i \mid C(0) = c_j) \\
 &= q_j(0) \cdot p_{i/j}(0), \quad i = 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w, \quad (56)
 \end{aligned}$$

where

$$q_j(0) = P(C(0) = c_j), \quad j = 1, 2, \dots, w, \quad (57)$$

are initial probabilities of the operation process  $C(t)$  defined in Chapter 4 and

$$\begin{aligned}
 p_{i/j}(0) &= P(Z(0) = z_i \mid C(0) = c_j), \quad (58) \\
 i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w,
 \end{aligned}$$

are conditional initial probabilities of the climate-weather change process  $Z(t)$  defined in Chapter 2 in case they are not conditional.

$$[pq_{ijkl}]_{w \times w} = \begin{bmatrix} pq_{1111} & pq_{1112} & \dots & pq_{111\nu} & pq_{1121} & pq_{1122} & \dots & pq_{112\nu} & \dots & pq_{11\nu 1} & pq_{11\nu 2} & \dots & pq_{11\nu w} \\ pq_{1211} & pq_{1212} & \dots & pq_{121\nu} & pq_{1221} & pq_{1222} & \dots & pq_{122\nu} & \dots & pq_{12\nu 1} & pq_{12\nu 2} & \dots & pq_{12\nu w} \\ \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots & \dots \\ pq_{w11} & pq_{w12} & \dots & pq_{w1\nu} & pq_{w21} & pq_{w22} & \dots & pq_{w2\nu} & \dots & pq_{w\nu 1} & pq_{w\nu 2} & \dots & pq_{w\nu w} \end{bmatrix} \quad (61)$$

Further, we introduce the probabilities

$$\begin{aligned}
 pq_{ijkl}, \quad i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w, \quad (59) \\
 k &= 1, 2, \dots, \nu, \quad l = 1, 2, \dots, w,
 \end{aligned}$$

of the transitions of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  between the operation states

$$\begin{aligned}
 zc_{ij} \text{ and } zc_{kl}, \quad i &= 1, 2, \dots, \nu, \quad (60) \\
 j &= 1, 2, \dots, w, \quad k = 1, 2, \dots, \nu, \quad l = 1, 2, \dots, w,
 \end{aligned}$$

and get their following matrix form

$$j = 1, 2, \dots, w, \quad k = 1, 2, \dots, \nu, \quad l = 1, 2, \dots, w,$$

where

$$q_{jl}, \quad j = 1, 2, \dots, w, \quad l = 1, 2, \dots, w, \quad (66)$$

In the case when the processes  $Z(t)$  and  $C(t)$  are dependent the probabilities of transitions between the operation states existing in (47) can be expressed either by

$$\begin{aligned}
 pq_{ijkl} &= p_{ik} \cdot q_{jl/ik} \quad i = 1, 2, \dots, \nu, \quad (62) \\
 j &= 1, 2, \dots, w, \quad k = 1, 2, \dots, \nu, \quad l = 1, 2, \dots, w,
 \end{aligned}$$

where

$$p_{ik}, \quad i = 1, 2, \dots, \nu, \quad k = 1, 2, \dots, \nu, \quad (63)$$

are transient probabilities of the operation process  $Z(t)$  defined in Chapter 2 and

$$\begin{aligned}
 q_{jl/ik}, \quad i &= 1, 2, \dots, \nu, \quad k = 1, 2, \dots, \nu, \quad (64) \\
 j &= 1, 2, \dots, w, \quad l = 1, 2, \dots, w,
 \end{aligned}$$

are conditional transient probabilities of the climate-weather change process  $C(t)$  defined in Chapter 4 in case they are not conditional or by

$$pq_{ijkl} = q_{jl} \cdot p_{ik/jl} \quad i = 1, 2, \dots, \nu, \quad (65)$$

are transient probabilities of the climate-weather change process  $C(t)$  defined in Chapter 4 and

$$\begin{aligned}
 p_{ik/jl}, \quad i &= 1, 2, \dots, \nu, \quad k = 1, 2, \dots, \nu, \quad (67) \\
 j &= 1, 2, \dots, w, \quad l = 1, 2, \dots, w,
 \end{aligned}$$

are conditional transient probabilities of the operation process  $Z(t)$  defined in Chapter 2 in case they are not conditional.

The matrix of conditional distribution functions

$$\begin{aligned}
 HC_{ijkl}(t) &= P(\theta C_{ijkl} < t), \quad t \in \langle 0, +\infty \rangle, \quad (68) \\
 i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w, \quad k = 1, 2, \dots, \nu, \\
 l &= 1, 2, \dots, w,
 \end{aligned}$$

of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  conditional sojourn times  $\theta C_{ijkl}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ ,  $k = 1, 2, \dots, \nu$ ,  $l = 1, 2, \dots, w$ , at the operation state  $zc_{ik}$ ,

$i=1,2,\dots,\nu, k=1,2,\dots,\nu$ , when the next operation following form  
state is  $zc_{jl}$ ,  $j=1,2,\dots,w, l=1,2,\dots,w$ , takes the

$$[HC_{ijkl}(t)]_{\nu \times \nu \times w}$$

$$= \begin{bmatrix} HC_{1111}(t) HC_{1112}(t) \dots HC_{111w}(t); HC_{1121}(t) HC_{1122}(t) \dots HC_{112w}(t); \dots; HC_{11v1}(t) HC_{11v2}(t) \dots HC_{11vw}(t) \\ HC_{1211}(t) HC_{1212}(t) \dots HC_{121w}(t); HC_{1221}(t) HC_{1222}(t) \dots HC_{122w}(t); \dots; HC_{12v1}(t) HC_{12v2}(t) \dots HC_{12vw}(t) \\ \dots \\ HC_{vw11}(t) HC_{vw12}(t) \dots HC_{vw1w}(t); HC_{vw21}(t) HC_{vw22}(t) \dots HC_{vw2w}(t); \dots; HC_{wv1}(t) HC_{wv2}(t) \dots HC_{wvw}(t) \end{bmatrix} \quad (69)$$

and the matrix of their corresponding conditional density functions

$$hc_{ijkl}(t) = \frac{d}{dt} [HC_{ijkl}(t)] \text{ for } t \in \langle 0, +\infty \rangle, i=1,2,\dots,\nu, j=1,2,\dots,w, k=1,2,\dots,\nu, l=1,2,\dots,w, \quad (70)$$

the form

$$[hc_{ijkl}(t)]_{\nu \times \nu \times w}$$

$$= \begin{bmatrix} hc_{1111}(t) hc_{1112}(t) \dots hc_{111w}(t); hc_{1121}(t) hc_{1122}(t) \dots hc_{112w}(t); \dots; hc_{11v1}(t) hc_{11v2}(t) \dots hc_{11vw}(t) \\ hc_{1211}(t) hc_{1212}(t) \dots hc_{121w}(t); hc_{1221}(t) hc_{1222}(t) \dots hc_{122w}(t); \dots; hc_{12v1}(t) hc_{12v2}(t) \dots hc_{12vw}(t) \\ \dots \\ hc_{vw11}(t) hc_{vw12}(t) \dots hc_{vw1w}(t); hc_{vw21}(t) hc_{vw22}(t) \dots hc_{vw2w}(t); \dots; hc_{wv1}(t) hc_{wv2}(t) \dots hc_{wvw}(t) \end{bmatrix}. \quad (71)$$

In the case when the critical infrastructure operation process  $Z(t)$  and climate-weather change process  $C(t)$  are dependent, the distribution functions existing in (70) can be expressed either by

$$HC_{ijk}(t) = P(\theta C_{ijk} < t) = P(\theta_{ik} < t \cap C_{jl} < t) \\ = H_{ik}(t) C_{jl/ik}(t), \quad t \in \langle 0, +\infty \rangle, \quad (72) \\ i=1,2,\dots,\nu, j=1,2,\dots,w, k=1,2,\dots,\nu, \\ l=1,2,\dots,w,$$

where

$$H_{ik}(t), \quad i=1,2,\dots,\nu, k=1,2,\dots,\nu, \quad (73)$$

are distribution functions defined of the sojourn lifetimes of the operation process  $Z(t)$  defined in Chapter 2 and

$$C_{jl/ik}(t) = P(C_{jl} < t \mid \theta_{ik} < t), \quad i=1,2,\dots,\nu, \quad (74) \\ j=1,2,\dots,w, k=1,2,\dots,\nu, l=1,2,\dots,w,$$

are conditional distributions of the sojourn lifetimes at the climate-weather states of the climate-weather

change process  $C(t)$  defined in Chapter 4 in case they are not conditional or by

$$HC_{ijk}(t) = P(\theta C_{ijk} < t) = P(\theta_{ik} < t \cap C_{jl} < t) \quad (75) \\ = C_{jl}(t) H_{ik/jl}(t), \quad t \in \langle 0, +\infty \rangle, \\ i=1,2,\dots,\nu, j=1,2,\dots,w, k=1,2,\dots,\nu, \\ l=1,2,\dots,w,$$

where

$$C_{jl}(t), \quad j=1,2,\dots,w, l=1,2,\dots,w, \quad (76)$$

are distribution functions defined of the sojourn lifetimes at the climate-weather states of the climate-weather change process  $C(t)$  defined in the Chapter 4 and

$$H_{ik/jl}(t) = P(\theta_{ik} < t \mid C_{jl} < t), \quad i=1,2,\dots,\nu, \quad (77) \\ j=1,2,\dots,w, k=1,2,\dots,\nu, l=1,2,\dots,w,$$

are conditional distributions of the sojourn lifetimes at the operation states of the critical infrastructure operation process  $Z(t)$  defined in Chapter 2 in case they are not conditional.

Hence, the density functions existing in (71) can be expressed either by

$$\begin{aligned} hc_{ijkl}(t) &= \frac{d}{dt}[HC_{ijkl}(t)] = \frac{d}{dt}[H_{ik}(t)C_{j/lk}(t)] \\ &= h_{ik}(t)C_{j/lk}(t) + H_{ik}(t)c_{j/lk}(t), \quad t \in \langle 0, +\infty \rangle, \quad (78) \\ i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w, \quad k = 1, 2, \dots, \nu, \\ l &= 1, 2, \dots, w, \end{aligned}$$

where

$$\begin{aligned} H_{ik}(t), \quad i &= 1, 2, \dots, \nu, \quad k = 1, 2, \dots, \nu, \quad (79) \\ \text{and } C_{j/lk}(t), \quad j &= 1, 2, \dots, w, \quad l = 1, 2, \dots, w, \end{aligned}$$

and

$$\begin{aligned} h_{ik}(t), \quad i &= 1, 2, \dots, \nu, \quad k = 1, 2, \dots, \nu, \quad \text{and } c_{j/lk}(t), \quad (80) \\ j &= 1, 2, \dots, w, \quad l = 1, 2, \dots, w, \end{aligned}$$

are respectively defined in Section 2.1 and Section 4.1 or by

$$\begin{aligned} hc_{ijkl}(t) &= \frac{d}{dt}[HC_{ijkl}(t)] = \frac{d}{dt}[C_{jl}(t)H_{ik/jl}(t)] \quad (81) \\ &= c_{jl}(t)H_{ik/jl}(t) + C_{jl}(t)h_{ik/jl}(t), \quad t \in \langle 0, +\infty \rangle, \\ i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w, \quad k = 1, 2, \dots, \nu, \\ l &= 1, 2, \dots, w, \end{aligned}$$

where

$$\begin{aligned} C_{jl}(t), \quad j &= 1, 2, \dots, w, \quad l = 1, 2, \dots, w, \quad (82) \\ \text{and } H_{ik/jl}(t), \quad i &= 1, 2, \dots, \nu, \quad k = 1, 2, \dots, \nu, \end{aligned}$$

and

$$\begin{aligned} c_{jl}(t), \quad j &= 1, 2, \dots, w, \quad l = 1, 2, \dots, w, \quad (83) \\ \text{and } h_{ik/jl}(t), \quad i &= 1, 2, \dots, \nu, \quad k = 1, 2, \dots, \nu, \end{aligned}$$

are respectively defined in Section 4.1 and Section 2.1 [EU-CIRCLE Report D2.1-GMU2, 2016].

We assume that the suitable and typical distributions suitable to describe the critical infrastructure operation process  $Z(t)$  conditional sojourn times  $\theta_{bl}$ ,  $b, l = 1, 2, \dots, \nu$ ,  $b \neq l$ , in the particular operation states are that defined in Section 4.2.4 [EU-CIRCLE Report D2.1-GMU2, 2016].

## 7. Critical infrastructure operation process related to climate-weather change process - prediction

Assuming that we have identified the unknown parameters of the critical infrastructure operation process related to climate-weather change  $ZC(t)$ ,  $t \in \langle 0, +\infty \rangle$ , that can take  $\nu w$ ,  $\nu, w \in N$ , different operation states  $zc_{11}, zc_{12}, \dots, zc_{\nu w}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ , defined in Section 5.6 [EU-CIRCLE Report D2.1-GMU2, 2016] and described by :

- the vector  $[pq_{ij}(0)]_{\nu w}$  of initial probabilities of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  staying at the initial moment  $t = 0$  at the operation and climate-weather states  $zc_{ij}$ ,  $i = 1, 2, \dots, \nu \in N$ ,  $j = 1, 2, \dots, w$ ;
- the matrix  $[pq_{ijkl}]_{\nu w \times \nu w}$  of the probabilities of transitions of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  between the operation states  $zc_{ij}$  and  $zc_{kl}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ ,  $k = 1, 2, \dots, \nu$ ,  $l = 1, 2, \dots, w$ ;
- the matrix  $[HC_{ijkl}(t)]_{\nu w \times \nu w}$  of the matrix of conditional distribution functions of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  conditional sojourn times  $\theta C_{ijkl}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ ,  $k = 1, 2, \dots, \nu$ ,  $l = 1, 2, \dots, w$ , at the operation state  $zc_{ik}$ ,  $i = 1, 2, \dots, \nu$ ,  $k = 1, 2, \dots, \nu$ , when the next operation state is  $zc_{jl}$ ,  $j = 1, 2, \dots, w$ ,  $l = 1, 2, \dots, w$ ,

we can predict this process basic characteristics.

### 7.1. Critical infrastructure operation process related to climate-weather change process characteristics – independent critical infrastructure operation process and climate-weather change process

The mean values of the conditional sojourn times  $\theta C_{ijkl}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ ,  $k = 1, 2, \dots, \nu$ ,  $l = 1, 2, \dots, w$ , at the operation state  $zc_{ik}$ ,  $i = 1, 2, \dots, \nu$ ,  $k = 1, 2, \dots, \nu$ , when the next operation state is  $zc_{jl}$ ,  $j = 1, 2, \dots, w$ ,  $l = 1, 2, \dots, w$ , are defined by [Kołowrocki, Soszyńska-Budny, 2011]

$$\begin{aligned}
 MN_{ijkl} &= E[\theta C_{ijkl}] = \int_0^{\infty} t dHC_{ijkl}(t) dt \\
 &= \int_0^{\infty} t h c_{ijkl}(t) dt, \quad (84) \\
 i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w, \quad k = 1, 2, \dots, \nu, \\
 l &= 1, 2, \dots, w.
 \end{aligned}$$

In the case when the processes  $Z(t)$  and  $C(t)$  are independent, according to (37) the expressions (84) take the form

$$\begin{aligned}
 MN_{ijkl} &= E[\theta C_{ijkl}] \\
 &= \int_0^{\infty} t [h_k(t) C_{jl}(t) + H_k(t) c_{jl}(t)] dt, \quad i = 1, 2, \dots, \nu, \quad (85) \\
 j &= 1, 2, \dots, w, \quad k = 1, 2, \dots, \nu, \quad l = 1, 2, \dots, w.
 \end{aligned}$$

Since from the formula for total probability, it follows that the unconditional distribution functions of the conditional sojourn times  $\theta C_{ij}$ , of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  at the operation states state

$$\begin{aligned}
 zc_{ij}, \quad i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w, \quad \text{are given by} \\
 HC_{ij}(t) &= \sum_{k=1}^{\nu} \sum_{l=1}^w p_{ijkl} HC_{ijkl}(t), \quad t \in \langle 0, +\infty \rangle, \quad (86) \\
 i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w,
 \end{aligned}$$

In the case when the processes  $Z(t)$  and  $C(t)$  are independent, according to (25) and (32) the expressions (76) take the form

$$\begin{aligned}
 HC_{ij}(t) &= \sum_{k=1}^{\nu} \sum_{l=1}^w p_{ik} q_{jl} H_{ik}(t) C_{jl}(t), \quad t \in \langle 0, +\infty \rangle, \quad (87) \\
 i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w,
 \end{aligned}$$

From (86) it follows that the mean values  $E[\theta C_{ij}]$  of the unconditional distribution functions of the conditional sojourn times  $\theta C_{ij}$ , of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  at the operation states  $zc_{ij}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ , are given by

$$\begin{aligned}
 MN_{ij} &= E[\theta C_{ij}] = \sum_{k=1}^{\nu} \sum_{l=1}^w p_{ijkl} MN_{ijkl}, \quad (88) \\
 i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w,
 \end{aligned}$$

where  $MN_{ijkl}$  are given by the formula (84).

In the case when the processes  $Z(t)$  and  $C(t)$  are independent, considering (87) and (32) the expression (78) takes the form

$$\begin{aligned}
 MN_{ij} &= E[\theta C_{ij}] = \sum_{k=1}^{\nu} \sum_{l=1}^w p_{ik} q_{jl} MN_{ijkl}, \quad (89) \\
 i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w,
 \end{aligned}$$

where  $MN_{ijkl}$  are given by the formula (85).

The transient probabilities of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  at the operation states  $zc_{ij}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ , can be defined by

$$\begin{aligned}
 pq_{ij}(t) &= P(ZC(t) = zc_{ij}), \quad t \in \langle 0, +\infty \rangle, \quad (90) \\
 i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w.
 \end{aligned}$$

In the case when the processes  $Z(t)$  and  $C(t)$  are independent the expression (90) for the transient probabilities can be expressed in the following way

$$\begin{aligned}
 pq_{ij}(t) &= P(ZC(t) = zc_{ij}) = P(Z(t) = z_i \cap C(t) = c_j) \\
 &= P(Z(t) = z_i) \cdot P(C(t) = c_j) = p_i(t) \cdot q_j(t), \quad (91) \\
 t &\in \langle 0, +\infty \rangle, \quad i = 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w,
 \end{aligned}$$

where

$$p_i(t) = P(Z(t) = z_i), \quad t \in \langle 0, +\infty \rangle, \quad i = 1, 2, \dots, \nu, \quad (92)$$

are the transient probabilities of the operation process  $Z(t)$  defined in Chapter 2 and

$$\begin{aligned}
 q_j(t) &= P(C(t) = c_j), \quad t \in \langle 0, +\infty \rangle, \quad (93) \\
 j &= 1, 2, \dots, w,
 \end{aligned}$$

are the transient probabilities of the climate-weather change process  $C(t)$  defined in Chapter 4.

The limit values of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  at the operation states  $zc_{ij}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ , can be found from [Kołowrocki, Soszyńska-Budny, 2011]

$$pq_{ij} = \lim_{t \rightarrow \infty} \frac{\pi_{ij} MN_{ij}}{\sum_{i=1}^{\nu} \sum_{j=1}^w \pi_{ij} MN_{ij}}, \quad i = 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w, \quad (94)$$

where  $MN_{ij}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ , are given by (89), while the steady probabilities  $\pi_{ij}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ , of the vector  $[\pi_{ij}]_{1 \times \nu w}$  satisfy the system of equations

$$\begin{cases} [\pi_{ij}] [pq_{ij}] = [\pi_{ij}] \\ \sum_{i=1}^{\nu} \sum_{j=1}^w \pi_{ij} = 1, \end{cases} \quad (95)$$

where  $p_{q_{ij}}, i=1,2,\dots,\nu, j=1,2,\dots,w, k=1,2,\dots,\nu, l=1,2,\dots,w$ , are given by (25).

In the case of a periodic system operation process, the limit transient probabilities  $p_{q_{ij}}, i=1,2,\dots,\nu, j=1,2,\dots,w$ , at the operation states given by (94), are the long term proportions of the critical infrastructure operation process  $ZC_{ij}(t)$  sojourn times at the particular operation states  $zc_{ij}, i=1,2,\dots,\nu, j=1,2,\dots,w$ .

Other interesting characteristics of the critical infrastructure operation process  $ZC_{ij}(t)$  possible to obtain are its total sojourn times  $\hat{\theta}_{C_{ij}}, i=1,2,\dots,\nu, j=1,2,\dots,w$ , at the particular operation states  $zc_{ij}, i=1,2,\dots,\nu, j=1,2,\dots,w$ , during the fixed system operation time. It is well known [Kołowrocki, Soszyńska-Budny, 2011] that the system operation process total sojourn times  $\hat{\theta}_{C_{ij}}$ , at the particular operation states  $zc_{ij}$ , for sufficiently large operation time  $\theta$ , have approximately normal distributions with the expected value given by

$$\hat{M}\hat{N}_{ij} = E[\hat{\theta}_{C_{ij}}] = p_{q_{ij}}\theta, \quad i=1,2,\dots,\nu, \quad j=1,2,\dots,w, \quad (96)$$

where  $p_{q_{ij}}, i=1,2,\dots,\nu, j=1,2,\dots,w$ , are given by (94).

## 7.2. Critical infrastructure operation process related to climate-weather change process characteristics – dependent critical infrastructure operation process and climate-weather change process

The mean values of the conditional sojourn times  $\theta_{C_{ijkl}}, i=1,2,\dots,\nu, j=1,2,\dots,w, k=1,2,\dots,\nu, l=1,2,\dots,w$ , at the operation state  $zc_{ik}, i=1,2,\dots,\nu, k=1,2,\dots,\nu$ , when the next operation state is  $zc_{jl}, j=1,2,\dots,w, l=1,2,\dots,w$ , are defined by [Kołowrocki, Soszyńska-Budny, 2011]

$$\begin{aligned} MN_{ijkl} &= E[\theta_{C_{ijkl}}] = \int_0^{\infty} t dHC_{ijkl}(t) dt \\ &= \int_0^{\infty} t hc_{ijkl}(t) dt, \quad i=1,2,\dots,\nu, \\ & \quad j=1,2,\dots,w, \quad k=1,2,\dots,\nu, \quad l=1,2,\dots,w. \end{aligned} \quad (97)$$

Since from the formula for total probability, it follows that the unconditional distribution functions of the conditional sojourn times  $\theta_{C_{ij}}$ , of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  at the operation states  $zc_{ij}, i=1,2,\dots,\nu, j=1,2,\dots,w$ , are given by

$$\begin{aligned} HC_{ij}(t) &= \sum_{k=1}^{\nu} \sum_{l=1}^w p_{ijkl} HC_{ijkl}(t), \quad t \in \langle 0, +\infty \rangle, \\ i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w, \end{aligned} \quad (98)$$

Hence, the mean values  $E[\theta_{C_{ij}}]$  of the unconditional distribution functions of the conditional sojourn times  $\theta_{C_{ij}}$ , of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  at the operation states  $zc_{ij}, i=1,2,\dots,\nu, j=1,2,\dots,w$ , are given by

$$\begin{aligned} MN_{ij} &= E[\theta_{C_{ij}}] = \sum_{k=1}^{\nu} \sum_{l=1}^w p_{ijkl} MN_{ijkl}, \\ i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w, \end{aligned} \quad (99)$$

where  $MN_{ijkl}$  are defined by the formula (87).

The transient probabilities of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  at the operation states  $zc_{ij}, i=1,2,\dots,\nu, j=1,2,\dots,w$ , can be defined by

$$\begin{aligned} p_{q_{ij}}(t) &= P(ZC(t) = zc_{ij}), \quad t \in \langle 0, +\infty \rangle, \\ i &= 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w. \end{aligned} \quad (100)$$

In the case when the processes  $Z(t)$  and  $C(t)$  are dependent the transient probabilities can be expressed either by

$$\begin{aligned} p_{q_{ij}}(t) &= P(ZC(t) = zc_{ij}) \\ &= P(Z(t) = z_i \cap C(t) = c_j) \\ &= P(Z(t) = z_i) \cdot P(C(t) = c_j | Z(t) = z_i) \\ &= p_i(t) \cdot q_{j|i}(t), \quad t \in \langle 0, +\infty \rangle, \quad i=1,2,\dots,\nu, \\ & \quad j=1,2,\dots,w, \end{aligned} \quad (101)$$

where

$$p_i(t) = P(Z(t) = z_i), \quad t \in \langle 0, +\infty \rangle, \quad i=1,2,\dots,\nu, \quad (102)$$



are transient probabilities of the operation process  $Z(t)$  defined in Chapter 2 and

$$q_{j/i}(t) = P(C(t) = c_j \mid Z(t) = z_i, t \in \langle 0, +\infty \rangle), \quad (103)$$

$$i = 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w,$$

are conditional transient probabilities of the climate-weather change process  $C(t)$  defined in Chapter 4 in case they are not conditional or by

$$pq_{ij}(t) = P(ZC(t) = zc_{ij}) = P(Z(t) = z_i \cap C(t) = c_j)$$

$$= P(C(t) = c_j) \cdot P(Z(t) = z_i \mid C(t) = c_j)$$

$$= q_j(t) \cdot p_{i/j}(t), \quad t \in \langle 0, +\infty \rangle, \quad i = 1, 2, \dots, \nu, \quad (104)$$

$$j = 1, 2, \dots, w,$$

where

$$q_j(t) = P(C(t) = c_j), \quad t \in \langle 0, +\infty \rangle, \quad (105)$$

$$j = 1, 2, \dots, w,$$

are transient probabilities of the operation process  $C(t)$  defined in Chapter 4 and

$$p_{i/j}(t) = P(Z(t) = z_i \mid C(t) = c_j), \quad t \in \langle 0, +\infty \rangle,$$

$$i = 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w,$$

are conditional transient probabilities of the climate-weather change process  $Z(t)$  defined in Chapter 2 in case they are not conditional.

The limit values of the critical infrastructure operation process related to climate-weather change  $ZC(t)$  at the operation states  $zc_{ij}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ , can be found from [Kołowrocki, Soszyńska-Budny, 2011]

$$pq_{ij} = \lim_{t \rightarrow \infty} \frac{\pi_{ij} MN_{ij}}{\sum_{i=1}^{\nu} \sum_{j=1}^w \pi_{ij} MN_{ij}}, \quad i = 1, 2, \dots, \nu, \quad j = 1, 2, \dots, w, \quad (106)$$

where  $MN_{ij}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ , are given by (99), while the steady probabilities  $\pi_{ij}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ , of the vector  $[\pi_{ij}]_{1, \nu \times w}$  satisfy the system of equations

$$\begin{cases} [\pi_{ij}][pq_{ij kl}] = [\pi_{ij}] \\ \sum_{i=1}^{\nu} \sum_{j=1}^w \pi_{ij} = 1. \end{cases} \quad (107)$$

In the case of a periodic system operation process, the limit transient probabilities  $pq_{ij}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ , at the operation states given by (107), are the long term proportions of the critical infrastructure operation process  $ZC_{ij}(t)$  sojourn times at the particular operation states  $zc_{ij}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ .

Other interesting characteristics of the critical infrastructure operation process  $ZC_{ij}(t)$  possible to obtain are its total sojourn times  $\hat{\theta}_{C_{ij}}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ , at the particular operation states  $zc_{ij}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ , during the fixed system operation time. It is well known [Kołowrocki, Soszyńska-Budny, 2011] that the system operation process total sojourn times  $\hat{\theta}_{C_{ij}}$ , at the particular operation states  $zc_{ij}$ , for sufficiently large operation time  $\theta$ , have approximately normal distributions with the expected value given by

$$M\hat{N}_{ij} = E[\hat{\theta}_{C_{ij}}] = pq_{ij} \theta, \quad i = 1, 2, \dots, \nu, \quad (108)$$

$$j = 1, 2, \dots, w,$$

where  $pq_{ij}$ ,  $i = 1, 2, \dots, \nu$ ,  $j = 1, 2, \dots, w$ , are given by (106).

## 8. Conclusions

The probabilistic model of the critical infrastructure operation process related to climate-weather change process presented in this Chapter is the basis for further considerations in particular tasks of the project. First, this model will be used to construct the integrated general safety probabilistic model of the critical infrastructure related to its operation process and climate-weather process [EU-CIRCLE Report D.e3.3-GMU3, CIOP Model5, 2016].

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